# The benefits of reducing health inequalities in 10 European populations 

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#### Abstract

In recent years there has been growing interest in studying the socio-economic inequalities in health. In this paper we use longitudinal data from SHARE survey to estimate the age and sex specific mortality rates by socioeconomic status (SES) for 10 European countries with the aim of studying the benefits of reducing mortality in the most disadvantaged classes.

First, we describe the existing inequalities by estimating the influence of SES (measured by two different proxies: the household total net worth and the education) on mortality between waves using Cox survival regression models. In a second step, we construct life tables for each combination of country, sex and SES, and we estimated the number of real deaths in the population. Then, some "inequality reduction" scenarios are depicted by reducing the SES gradient for each country and providing an estimate of the hypothetical saved life-years.


## 1. Introduction

In recent years there has been growing interest in studying the socio-economic inequalities in health in many European countries (Mackenbach et al., 2007; Avendano et al., 2009). Reducing these health inequalities has become an important policy objective. It is widely accepted, indeed, that socioeconomic inequalities in health are an unfair feature of Western societies, as health is good that all citizens - regardless of their socioeconomic status - should equally get access to.

In this paper we want to provide estimates for health disadvantage in various European countries, using survey data. In particular, we will use data from SHARE (Survey of Health, Ageing and Retirement in Europe) surveys. SHARE provides us with longitudinal information on people aged over 50 . We therefore estimate the age and sex specific mortality rates by socioeconomic status for all the available country. In addition, we estimate the benefits of reducing mortality in the most disadvantaged classes, depicting some "inequality reduction" scenarios obtained by reducing the SES gradient for each country and by providing an estimate of the hypothetical saved life-years.

## 2. Data and methods

### 2.1 Data

Data from the Survey of Health, Ageing and Retirement in Europe (SHARE) are used. The survey is a panel database providing information on health and socio-economic status of noninstituzionalized adults aged 50 or over ${ }^{1}$ representing the various European regions (Börsch-Supan et al., 2005). In this way comparable information across countries are available. In particular, in the 2004 SHARE baseline study representative samples were obtained for eleven countries which are the focus of our paper ${ }^{2}$ : Denmark and Sweden (representing Scandinavian countries), Austria, Belgium, France, Germany, and Netherlands (representing the Central Europe), Greece, Italy, and Spain (for the Mediterranean area). The second wave of data collection was conducted in 20062007 and the third one in 2008-2009.

We used information on the socio-economic status (SES) of individuals in the first wave and we considered whether the same individuals are alive in the following waves. For dead individuals the

1 The focus only on population aged 50 or over is not a limitation since most of mortality is concentrated on ages over 50. In fact, a limitation may be the fact that only non-instituzionalized individuals are considered and clearly the most healthy: as a consequence the mortality may be underestimated.
2 Further data were collected in Israel in 2005-2006 and from the second wave (in 2006-2007) also Poland and the Czech Republic joined SHARE. These three countries are not used in this paper. Switzerland is also excluded in the analyses.
date of death is available so that we can consider the socio-economic status as a determinant of individuals' survival.

SHARE allows us to use different indicators of socioeconomic status.
Following the definition used by other researches (see Avendano et al., 2009), the first indicator that we consider is the household total net worth. Following Avendano et al. (2009) this is "the sum of all financial (net stock value, mutual funds, bonds, and savings) and housing wealth (value of primary residence net of mortgage, other real estate value, own business share, and owned cars) minus liabilities". Missing items were imputed using the methodology of multiple imputation (see SHARE Release Guide 2.5 .0 waves 1\& 2, Mannheim Research Institute for the Economics of Aging, 2011). The differences in the number of household members are accounted dividing wealth by the square root of household size (Buhmann et al., 1998; Huisman et al 2003; Avendano et al., 2009). In the following analyses, we collapsed wealth into country specific quintiles.

The second indicator of socio-economic status is education. In the survey it is measured using the ISCED (International Standard Classification of Education) coding; then we grouped the different levels into three categories: low corresponding to the ISCED-codes from 0 to 2 (lower secondary school or lower), medium corresponding to the ISCED-code 3 (upper secondary school), and high including ISCED-codes from 4 to 6 (postsecondary).

### 2.2 Methods

Our analysis of health inequalities and potential scenarios of their reductions consisted in three steps.

First, we start with the accurate description of existing inequalities. In particular, we estimate the influence of SES on mortality by mean of Cox survival regression models. Net of age and sex, they estimate the effects of SES on the risk of death considering the first wave as a starting time.

In a second step, from the results of the regression models we construct life tables for each combination of country, sex and SES status. Predicted values of mortality rates have been obtained by the estimated models and from these predicted values we constructed the life tables. From the life tables we take five-years age-specific mortality rates by SES and referring to the population by gender, SES and countries (obtained from weighted survey samples), we estimated the number of real deaths in the population.
Then, considering separately men and women, some "inequality reduction" scenarios are depicted by reducing the SES gradient for each country and providing an estimate of the saved life-years (clearly, all scenarios are hypothetical).

## 3. Empirical analysis

### 3.1 Inequalities based on wealth as a SES proxy

First of all, table 1 presents for each gender and for each country the hazard ratios of SES in the regression models.

Table 1. Estimated hazard ratios for wealth quintiles of Cox regression models net of age and sex for the different countries.

|  | 1st quintile | 2nd quintile | 3rd quintile | 4th quintile | 5th quintile |
| :--- | :---: | :---: | :---: | :---: | :---: |
| AUSTRIA | 1.327 | 0.786 | 0.918 | 0.751 | 1.000 (ref) |
| BELGIUM | 1.179 | 1.160 | 1.035 | 0.896 | 1.000 (ref) |
| DENMARK | 2.051 | 2.064 | 1.517 | 0.959 | 1.000 (ref) |
| FRANCE | 1.602 | 1.796 | 1.046 | 0.634 | 1.000 (ref) |
| GERMANY | 1.169 | 1.205 | 0.997 | 0.769 | 1.000 (ref) |
| GREECE | 3.518 | 2.820 | 2.023 | 1.944 | 1.000 (ref) |
| ITALY | 1.307 | 1.522 | 1.171 | 1.152 | 1.000 (ref) |
| NETHERLANDS | 2.232 | 2.259 | 1.644 | 2.040 | 1.00 (ref) |
| SPAIN | 1.490 | 1.176 | 1.051 | 0.690 | 1.00 (ref) |
| SWEDEN | 2.310 | 1.250 | 1.108 | 1.184 | 1.00 (ref) |

Table 2 describes the existing inequalities for men and women of the different countries considering as a synthetic measure of mortality the life expectancy at the age of 50. It should be noted that these life expectancies are constantly higher than those reported by official statistics. For example, France life expectancy at 50 reported by the national institute of statistics (INSEE) is 29.09 for men and 34.96 for women, while life expectancies at 50 reported in table 2 are all higher than these values. This discrepancy is certainly due to the fact that all individuals in institutions (included hospitals) are not included in the SHARE sample. Moreover, it is likely that individuals living at home but with severe health conditions have not participated to the survey. Therefore, we should expect that individuals of SHARE sample have a better health - and, consequently a higher life expectancy that the whole population. We should keep this in mind when commenting the results of our computations.
Both the hazard ratios and the estimated life expectancies at 50 by wealth quintiles reveal a varying level of inequality in each country. In Greece, for example the mortality rate of men aged over 50 belonging to the $1^{\text {st }}$ wealth quintile (i.e. the poorest group) increases of $251 \%$ with respect men belonging to the $5^{\text {th }}$ wealth quintile (i.e. the richest). This brings about a difference of about 9 years of life expectancy. In other countries (for example in Germany or in Belgium) the difference across the wealth quintiles is milder. It should also be noted that not in every country we find the highest mortality among the poorest and the lowest among the richest. In Austria, for instance, the highest value of life expectancy is found for the $4^{\text {th }}$ wealth quintile. Generally speaking, we can say that
there is an increasing mortality between the richest population and the poorest, and in all countries the life expectancy among the $1^{\text {st }}$ quintile group is lower than the life expectancy of the $5^{\text {th }}$ quintile group. However such a decrease is not linear, and it is often found that a quintile group has a higher life expectancy than the adjacent richer one.

Table 2. Estimated life expectancy at 50 by wealth quintiles, sex and countries.

| MEN |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | 1st quintile | 2nd quintile | 3rd quintile | 4th quintile | 5th quintile |
| AUSTRIA | 31.351 | 36.478 | 35.024 | 36.765 | 34.199 |
| BELGIUM | 35.070 | 35.250 | 36.247 | 37.382 | 36.451 |
| DENMARK | 30.323 | 30.303 | 33.068 | 37.209 | 36.867 |
| FRANCE | 31.305 | 30.337 | 34.897 | 38.949 | 35.304 |
| GERMANY | 32.830 | 32.532 | 34.019 | 35.963 | 34.015 |
| GREECE | 35.253 | 37.046 | 39.572 | 39.899 | 44.055 |
| ITALY | 28.784 | 27.372 | 29.740 | 29.882 | 31.274 |
| NETHERLANDS | 32.103 | 32.013 | 34.568 | 32.854 | 38.359 |
| SPAIN | 27.294 | 29.470 | 30.523 | 34.379 | 30.950 |
| SWEDEN | 30.219 | 35.060 | 36.018 | 35.455 | 36.815 |
|  |  |  |  |  |  |
|  |  |  |  |  |  |
| AUSTRIA | 1 st quintile | 2 nd quintile | 3 3rd quintile | 4 th quintile | 5 th quintile |
| BELGIUM | 35.024 | 39.524 | 38.257 | 39.981 | 37.505 |
| DENMARK | 41.671 | 41.749 | 42.448 | 43.354 | 42.689 |
| FRANCE | 34.185 | 34.126 | 36.852 | 40.543 | 40.221 |
| GERMANY | 36.814 | 35.867 | 40.060 | 43.271 | 40.413 |
| GREECE | 38.100 | 37.894 | 39.304 | 41.064 | 39.269 |
| ITALY | 37.863 | 39.515 | 41.750 | 42.041 | 45.612 |
| NETHERLANDS | 33.970 | 32.573 | 34.937 | 35.125 | 36.371 |
| SPAIN | 36.854 | 36.779 | 39.147 | 37.563 | 42.463 |
| SWEDEN | 31.574 | 33.747 | 34.784 | 38.373 | 35.198 |

### 3.2 Health inequalities reduction scenarios based on wealth as a SES proxy

We use age-specific mortality rates referring to 5 -year age groups ( $50-54,55-59, \ldots 85+$ ) by wealth quintiles obtained from Cox regression models and we multiply these mortality rates by the population at risk by wealth quintiles. In this way, we obtain an estimated number of deaths, by age groups and wealth quintiles for each country.

Subsequently, we simulate the number of life-years that would be gained if people of lower SES experienced the lower mortality rates of those of higher SES.

In particular, we considered three different scenarios:

1. mortality rates of the $1^{\text {st }}$ wealth quintile decrease to those of the $2^{\text {nd }}$;
2. mortality rates of the $1^{\text {st }}$ and $2^{\text {nd }}$ wealth quintile decrease to those of the $3^{\text {rd }}$;
3. all individuals have the mortality rates of the $3^{\text {rd }}$ quintile.

Scenarios 1 to 2 follow a successively more ambitious order, with scenario 2 being more ambitious than the first one. The idea of the third scenario is to impose zero costs - health losses to higher groups would exactly offset health gains to lower groups.

If we had a monotonic association between wealth and mortality, scenarios 1 and 2 should provide a reduction of the number of expected deaths, but since in many countries the relationship is not monotonic (as shown by Tables 1 and 2) we will found cases in which the number of expected deaths increases. In Denmark, for instance, since the life expectancy of the $1^{\text {st }}$ quintile group is higher than that of the $2^{\text {nd }}$ group, we should expect a increased number of deaths for Scenario 1.

By comparing the number of deaths simulated in the different scenarios to the number of deaths in the initial situation (Table 3), we can derive the number of deaths saved in each scenario.
These estimates are reported in Table 4.

Table 3. Estimated number of deaths by wealth quintiles, sex and countries.

| MEN |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | 1st quintile | 2nd quintile | 3rd quintile | 4th quintile | 5th quintile |
| AUSTRIA | 20,213 | 13,724 | 11,781 | 12,730 | 15,141 |
| BELGIUM | 22,151 | 18,087 | 16,423 | 12,822 | 15,225 |
| DENMARK | 15,698 | 15,233 | 12,817 | 7,374 | 6,947 |
| FRANCE | 148,870 | 168,072 | 117,583 | 68,688 | 109,814 |
| GERMANY | 202,100 | 162,814 | 149,262 | 116,351 | 152,225 |
| GREECE | 24,243 | 20,335 | 15,369 | 13,522 | 4,797 |
| ITALY | 260,369 | 215,262 | 143,406 | 201,996 | 177,525 |
| NETHERLANDS | 52,520 | 32,995 | 15,911 | 23,235 | 15,700 |
| SPAIN | 162,662 | 132,485 | 134,349 | 80,239 | 98,884 |
| SWEDEN | 32,897 | 22,763 | 18,055 | 19,206 | 12,766 |
|  |  |  |  |  |  |
|  | 1 st quintile | 2 nd quintile | 3 rd quintile | 4 th quintile | 5 th quintile |
| AUSTRIA | 42,284 | 14,388 | 16,168 | 14,253 | 14,062 |
| BELGIUM | 27,780 | 19,420 | 11,289 | 15,554 | 13,643 |
| DENMARK | 26,721 | 20,431 | 12,654 | 6,610 | 6,163 |
| FRANCE | 282,779 | 171,474 | 107,745 | 61,505 | 101,800 |
| GERMANY | 535,144 | 211,382 | 95,917 | 78,265 | 82,585 |
| GREECE | 34,922 | 21,533 | 12,989 | 11,914 | 3,920 |
| ITALY | 371,767 | 215,970 | 155,934 | 180,119 | 144,115 |
| NETHERLANDS | 75,536 | 46,333 | 12,864 | 28,004 | 9,014 |
| SPAIN | 192,651 | 121,447 | 123,436 | 92,784 | 144,139 |
| SWEDEN | 60,343 | 22,935 | 15,019 | 15,893 | 7,533 |

Generally speaking, among these three scenarios, the second one provides the highest reduction of deaths, but the situation varies across countries. In Austria, for example, scenario 1 provides a much higher reduction of deaths than scenario 2 . All scenarios in at least one case do not generate a reduction of deaths but an increase. This happens for scenario 1 in several countries, for scenario 3 in Denmark.

We then have to take into account the fact that those individuals whose lives would be saved in 2004 would be expected to live many more years beyond 2004, on average. To do so, we consider the life expectancies by 5 -years age groups for each of the SES classes. The total number of life years saved with improved mortality is equal to the number of lives saved in 2004 multiplied by remaining life expectancy, for each age group and SES class. Table 5 reports these data.

Table 4. Estimated number of individual whose lives would be saved under alternative scenarios by sex and countries.

|  | MEN |  |  |
| :--- | :---: | :---: | :---: |
|  | Scenario 1 | Scenario 2 | Scenario 3 |
| AUSTRIA | 6,443 | 3,052 | 2,070 |
| BELGIUM | 335 | 3,722 | 1,993 |
| DENMARK | -29 | 6,118 | -503 |
| FRANCE | $-12,452$ | 96,742 | 56,480 |
| GERMANY | $-5,246$ | 46,198 | 16,812 |
| GREECE | 3,551 | 12,148 | 7,906 |
| ITALY | $-30,615$ | 63,029 | 36,605 |
| NETHERLANDS | -352 | 17,640 | 14,225 |
| SPAIN | 26,621 | 49,249 | 13,674 |
| SWEDEN | 11,002 | 14,672 | 14,572 |
|  |  |  |  |
|  | Scenario 1 | Scenario 2 | Scenario 3 |
| AUSTRIA | 11,312 | 6,755 | 5,219 |
| BELGIUM | 185 | 2,852 | 1,300 |
| DENMARK | -107 | 8,121 | 3,182 |
| FRANCE | $-19,410$ | 107,582 | 79,098 |
| GERMANY | $-9,606$ | 70,908 | 52,980 |
| GREECE | 4,175 | 13,899 | 11,091 |
| ITALY | $-38,494$ | 61,018 | 40,631 |
| NETHERLANDS | -441 | 20,377 | 19,501 |
| SPAIN | 29,842 | 51,637 | 16,041 |
| SWEDEN | 16,538 | 20,280 | 20,382 |

* a negative number indicates that the number of deaths under that scenario is higher than that observed in real data.

The increase in deaths observed in Table 4 is reflected in results of Table 5. In fact, in some cases, despite a reduction of deaths, a negative number of life years was saved (in other words, a reduction of years saved): this is the case of scenario 3 for men in Denmark, for example. This might seem odd, but it depends on the fact that the number of saved lives is not uniformly distributed over the age groups. In this case, we have an increase of deaths in the younger age groups - for which life expectancy is higher - and a reduction for the older groups. Therefore the number of deaths increased in the first groups accounts for a higher number of life years lost than the number of deaths reduced in the older groups.

Table 5. Total number of life years saved under alternative scenarios by sex and countries.

|  | MEN |  |  |
| :--- | :---: | :---: | :---: |
|  | Scenario 1 | Scenario 2 | Scenario 3 |
| AUSTRIA | 82,432 | 33,912 | 22,114 |
| BELGIUM | 3,602 | 46,702 | 24,970 |
| DENMARK | -126 | 67,719 | $-18,748$ |
| FRANCE | $-106,746$ | $1,080,192$ | 619,732 |
| GERMANY | $-55,800$ | 471,245 | 191,215 |
| GREECE | 40,294 | 152,680 | 81,480 |
| ITALY | $-246,280$ | 559,343 | 281,790 |
| NETHERLANDS | $-2,394$ | 172,366 | 147,007 |
| SPAIN | 245,804 | 481,219 | 125,364 |
| SWEDEN | 103,597 | 144,712 | 141,910 |
|  |  |  |  |
|  | Wcenario 1 | Scenario 2 | Scenario 3 |
| AUSTRIA | 141,058 | 65,754 | 56,108 |
| BELGIUM | 1,992 | 33,849 | 16,557 |
| DENMARK | -987 | 77,127 | 18,754 |
| FRANCE | $-178,985$ | $1,011,561$ | 819,064 |
| GERMANY | $-77,013$ | 568,329 | 457,754 |
| GREECE | 45,669 | 159,706 | 128,839 |
| ITALY | $-325,702$ | 586,785 | 411,509 |
| NETHERLANDS | $-3,065$ | 163,838 | 176,102 |
| SPAIN | 283,683 | 481,684 | 138,046 |
| SWEDEN | 151,065 | 146,969 | 194,263 |

### 3.3 Inequalities based on education as SES proxy

A similar approach can be followed using education as the SES proxy.
Cox survival regression models are used with education, sex and age as covariates (Table 6 reports the hazard ratios estimates for education) to estimate age-specific mortality rates and life expectancies (Table 7 reports the life expectancies at the age of 50 ).

Table 6. Estimated hazard ratios for educational levels of Cox regression models net of age and sex for the different countries.

|  | Low | Medium | High |
| :--- | :---: | :---: | :---: |
| AUSTRIA | 2.873 | 2.208 | 1.000 (ref) |
| BELGIUM | 1.378 | 1.252 | 1.000 (ref) |
| DENMARK | 1.256 | 1.354 | 1.000 (ref) |
| FRANCE | 5.138 | 2.450 | 1.000 (ref) |
| GERMANY | 1.260 | 1.208 | 1.000 (ref) |
| GREECE | 1.361 | 1.024 | 1.000 (ref) |
| ITALY | 0.897 | 0.494 | 1.000 (ref) |
| NETHERLANDS | 1.092 | 0.874 | 1.00 (ref) |
| SPAIN | 0.992 | 0.556 | 1.00 (ref) |
| SWEDEN | 1.611 | 2.194 | 1.00 (ref) |

Also these tables, as the corresponding ones obtained considering wealth as a SES proxy, reveal a varying level of inequality in each country. We need, however, to be cautious in interpreting the
results reported. These are particularly odd for Italy and Spain, where education seems to increase mortality rather than reduce it, in contradiction with most of the existing literature. It should be noted that the proportion of high educated individuals (the reference group) is very low in Italy and Spain, so the strange effect of education might partly depend on this.

Table 7. Estimated life expectancy at 50 by educational levels, sex and countries.

|  | MEN |  |  |
| :--- | :---: | :---: | :---: |
|  | Low | Medium | High |
| AUSTRIA | 31.516 | 34.092 | 40.968 |
| BELGIUM | 35.186 | 36.026 | 37.935 |
| DENMARK | 33.045 | 32.345 | 35.085 |
| FRANCE | 30.666 | 37.074 | 43.390 |
| GERMANY | 33.056 | 33.462 | 34.843 |
| GREECE | 37.865 | 39.86 | 40.018 |
| ITALY | 29.082 | 34.553 | 28.062 |
| NETHERLANDS | 32.967 | 34.699 | 33.649 |
| SPAIN | 30.338 | 35.326 | 30.247 |
| SWEDEN | 33.996 | 31.717 | 37.562 |
|  | WOMEN |  |  |
|  | Low |  |  |
| AUSTRIA | 35.928 | Medium | High |
| BELGIUM | 41.822 | 42.262 | 43.889 |
| DENMARK | 35.889 | 35.255 | 43.752 |
| FRANCE | 37.371 | 42.711 | 47.839 |
| GERMANY | 38.287 | 38.553 | 39.938 |
| GREECE | 39.809 | 41.659 | 41.902 |
| ITALY | 33.96 | 39.073 | 32.949 |
| NETHERLANDS | 37.344 | 38.952 | 38.018 |
| SPAIN | 34.379 | 39.319 | 34.277 |
| SWEDEN | 38.464 | 36.307 | 41.693 |

### 3.4 Health inequalities reduction scenarios based on education as a SES proxy

The number of deaths is obtained multiplying age-specific mortality rates for education groups by the population at risk (Table 8).

Following in principle the approach used above, we can simulate the number of life-years that would be gained if people of lower educational groups experienced the lower mortality rates of those of higher educational levels. Three different scenarios are considered:

1. mortality rates of individual with low education decrease to those of individuals with a medium educational level.
2. all individuals have the mortality rates of the higher educated ones;
3. all individuals have the mortality rates of the individuals with a medium educational level.

Table 8. Estimated number of deaths by educational levels, sex and countries.

|  | Low | Medium | High |
| :--- | :---: | :---: | :---: |
| AUSTRIA | 17.595 | 31.680 | 17.645 |
| BELGIUM | 49.140 | 19.551 | 16.530 |
| DENMARK | 19.125 | 24.580 | 14.450 |
| FRANCE | 500.469 | 100.396 | 39.366 |
| GERMANY | 87.172 | 470.230 | 222.176 |
| GREECE | 62.987 | 10.118 | 5.267 |
| ITALY | 645.874 | 215.826 | 83.356 |
| NETHERLANDS | 82.972 | 29.430 | 28.248 |
| SPAIN | 532.421 | 19.071 | 41.025 |
| SWEDEN | 73.433 | 14.954 | 18.082 |
|  |  |  |  |
|  | Low | WOMEN |  |
| AUSTRIA | 61.395 | 30.979 | 7.725 |
| BELGIUM | 58.250 | 17.229 | 12.577 |
| DENMARK | 42.932 | 22.540 | 8.836 |
| FRANCE | 307.999 | 342.798 | 15.844 |
| GERMANY | 580.875 | 335.300 | 97.517 |
| GREECE | 67.814 | 13.664 | 3.002 |
| ITALY | 1.018 .123 | 38.511 | 23.138 |
| NETHERLANDS | 133.283 | 23.041 | 16.867 |
| SPAIN | 631.818 | 9.670 | 31.589 |
| SWEDEN | 99.361 | 12.174 | 11.184 |

Table 9 reports the estimates of the number of deaths saved in each scenario, obtained comparing the number of deaths simulated in the different scenarios to the number of deaths in the initial situation (of Table 8).

Once again, Scenario 2 looks as the most ambitious, as it provides the highest number of lives "saved" (with the exception of Italy and Spain because the above mentioned strange effect of education on mortality in these countries, and of Netherlands).
Table 10 reports the total number of life years saved with improved mortality under the different scenarios.

Table 9. Estimated number of individual whose lives would be saved under alternative scenarios by sex and countries.

|  | MEN |  |  |
| :--- | :---: | :---: | :---: |
|  | Scenario 1 | Scenario 2 | Scenario 3 |
| AUSTRIA | 3.843 | 27.143 | $-7,077$ |
| BELGIUM | 3.727 | 14.359 | 692 |
| DENMARK | -1.054 | 8.366 | $-5,053$ |
| FRANCE | 214.328 | 391.556 | 179,506 |
| GERMANY | 3.316 | 78.715 | $-29,505$ |
| GREECE | 10.856 | 11.734 | 10,776 |
| ITALY | 273.752 | -166.464 | 311,365 |
| NETHERLANDS | 13.397 | 2.353 | 16,277 |
| SPAIN | 191.650 | -16.647 | 209,320 |
| SWEDEN | -17.966 | 27.743 | $-30,595$ |
| WOMEN |  |  |  |
|  |  |  |  |
| AUSTRIA | Scenario 1 | Scenario 2 | Scenario 3 |
| BELGIUM | 9.320 | 39.883 | 4,186 |
| DENMARK | 3.020 | 11.079 | 1,473 |
| FRANCE | -1.922 | 9.661 | $-4,227$ |
| GERMANY | 159.631 | 304.985 | 145,913 |
| GREECE | 13.268 | 114.995 | -733 |
| ITALY | 10.079 | 11.742 | 9,967 |
| NETHERLANDS | 324.970 | -114.828 | 336,525 |
| SPAIN | 17.327 | 5.362 | 18,435 |
| SWEDEN | 200.427 | -11.729 | 209,743 |

* a negative number indicates that the number of deaths under that scenario is higher than that observed in real data.

Table 10. Total number of life years saved under alternative scenarios by sex and countries.

|  | MEN |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | Scenario 1 | Scenario 2 | Scenario 3 |  |  |
| AUSTRIA | 50.668 | 495.547 | $-96,575$ |  |  |
| BELGIUM | 45.599 | 190.835 | 3,569 |  |  |
| DENMARK | -8.477 | 106.432 | $-54,352$ |  |  |
| FRANCE | 2.535 .021 | 6.173 .183 | $2,037,912$ |  |  |
| GERMANY | 31.929 | 788.208 | $-289,936$ |  |  |
| GREECE | 126.902 | 138.048 | 125,659 |  |  |
| ITALY | 3.478 .302 | -1.385 .444 | $3,998,139$ |  |  |
| NETHERLANDS | 137.900 | 21.481 | 170,097 |  |  |
| SPAIN | 2.335 .931 | -211.213 | $2,620,868$ |  |  |
| SWEDEN | -122.770 | 289.350 | $-225,204$ |  |  |
|  | WOMEN |  |  |  |  |
|  | Scenario 1 | Scenario 2 | Scenario 3 |  |  |
| AUSTRIA | 118.216 | 643.371 | 38,221 |  |  |
| BELGIUM | 40.983 | 156.885 | 17,373 |  |  |
| DENMARK | -17.008 | 111.651 | $-50,304$ |  |  |
| FRANCE | 2.410 .432 | 5.118 .601 | $2,143,051$ |  |  |
| GERMANY | 97.423 | 1.194 .207 | $-76,090$ |  |  |
| GREECE | 138.049 | 165.496 | 135,396 |  |  |
| ITALY | 4.031 .140 | -1.251 .323 | $4,282,652$ |  |  |
| NETHERLANDS | 177.261 | 51.940 | 189,986 |  |  |
| SPAIN | 2.558 .004 | -161.087 | $2,712,657$ |  |  |
| SWEDEN | -130.019 | 282.586 | $-213,138$ |  |  |

## 5. Future developments

In the future, we intend to define other scenarios, also in the light of the results provided by those we have depicted above. A scenario we can take into consideration is a refinement of scenario 3: similarly to scenario 3 , it also pivots the social gradient about the level of the intermediate class, but only $50 \%$ of the way to becoming a horizontal line. In practice, this is achieved by halving the coefficients of the Cox regression models. This looks like as the least implausible but preliminary analyses showed that - especially when we use wealth as SES indicator - it provides an increase of the number of deaths in too many countries. In some cases this might depend on the fact that halving the Cox regression coefficients we actually half the $\log$ of the hazard ratio. Therefore, the reduction of higher SES groups mortality rates is greater than the increase of lower SES groups mortality rates. We can therefore provide a different scenario in which the hazard ratios and not the Cox coefficients are halved. In this way we expect to find fewer countries with a negative reduction of the number of deaths. A further refinement could be provided by assuming that mortality of all SES groups will decrease but with a decreasing rate over wealth quintiles. This looks a more likely scenario, as it would be difficult to believe that mortality decreases only for low SES groups. Finally, we can decide to change the mortality rates of countries more gradually. Since we have been using five-years mortality rates, this can be done projecting the over-50 population for ten, rather than five, years. In this way, in the first five years we can apply an intermediate set of mortality rates, which will basically be an average of the original rates and the rates defined by the scenarios. In the next five years, the rates defined by the scenario would be applied.

In addition, another future step consists on providing an estimate of the monetary expected benefits due to inequality reduction for each country and for each scenario, basing on available estimates of the value of a statistical life in each country. Again, the two measures of SES (the household total net worth and education) will be used alternatively.
Lastly, future work intends to examine also the benefits of reducing morbidity: a cross-sectional approach with data from the survey HITT (Health in Time of Transitions) will be used for this aim.

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